

# Foundations of Causal Discovery

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KDD Causality Workshop 2016

### data sample



#### assumptions, e.g.

- causal Markov
- causal faithfulness
- functional form
- etc.



### data sample inference algorithm

➟



#### assumptions, e.g.

- causal Markov
- causal faithfulness
- functional form
- etc.

### equivalence classes









➟







#### equivalence classes



### model specifications







equivalence classes





#### data sample



#### equivalence classes





statistical





*x y z w* truth (unknown) data sample *w x y z* samples ?

inference

statistical<br>inference

statistical

 $x \perp y \mid \{z, w\}$ graphical connection

conditions • Markov • faithfulness

 $x \perp y | \{z, w\}$ probabilistic independence



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### Causal Markov

x is independent of its non-descendents given its parents in the causal graph



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Violations of Causal Markov

- quantum mechanics
- [unmeasured common causes]
- [mixtures of populations]
- [variables are not distinct, or too coarsely grained]

*x y z*  $w$ *v*  $u$ 

### Causal Faithfulness

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### Violations of Causal Faithfulness

- canceling pathways
- matching pennies cases
- [small sample sizes and near violations of faithfulness]





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All graphs in an equivalence class have:

- same adjacencies ("skeleton")
- same unshielded colliders

[Verma & Pearl 1990, Frydenberg 1990]

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[Verma & Pearl 1990, Frydenberg 1990]



### assumptions assumptions • Markov

- faithfulness
- acyclicity
- causal sufficiency



### equivalence class

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- same unshielded colliders






















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- weaken faithfulness

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- **Include more general data collection set-ups** (and see how assumptions can be adjusted and what equivalence class results)
	- experimental evidence
	- multiple (overlapping) data sets
	- relational data

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## Limitations

*For linear Gaussian and for multinomial causal relations, an algorithm that identifies the Markov equivalence class of the true model is complete.* (Pearl & Geiger 1988, Meek 1995)

# Linear non-Gaussian method (LiNGaM)

• Linear causal relations:

$$
x_i = \sum_{x_j \in \mathbf{Pa}(x_i)} \beta_{ij} x_j + \epsilon_j
$$

- Assumptions:
	- causal Markov
	- causal sufficiency
	- acyclicity

[Shimizu et al., 2006]

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 $\blacktriangleright$  If  $\epsilon_j \sim$  non-Gaussian, then the true graph is **uniquely identifiable** from the joint distribution.

[Shimizu et al., 2006]

True model

$$
y = \beta x + \epsilon_y
$$



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 $x \perp\!\!\!\perp \epsilon_y$ 

True model

$$
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Backwards model

$$
x = \theta y + \tilde{\epsilon}_x
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 $\epsilon_x$ 

 $\epsilon_y$ 





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True model

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 $(x)$   $\leftarrow$   $(y)$  $\tilde{\epsilon}_x$   $\tilde{\epsilon}_y$ 

 $\epsilon_y$ 

 $\boldsymbol{x}$ 

 $\epsilon_x$ 

 $x \perp \!\!\!\perp \epsilon_y$ 

 $y \perp \tilde{\epsilon}_x$ 

Backwards model

$$
x = \theta y + \tilde{\epsilon}_x
$$

 $\tilde{\epsilon}_x = x - \theta y$  $= x - \theta(\beta x + \epsilon_y)$  $= (1 - \theta \beta)x - \theta \epsilon_y$ 

#### True model





#### Backwards model

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 $y \perp \tilde{\epsilon}_x$ 



# Why Normals are unusual

For backwards model





### Why Normals are unusual

 $y = \beta x + \epsilon_y$ Forwards model For backwards model ✏*y x* )  $\cdot$  *(y*  $\epsilon$ <sub>*x*</sub> ?  $\widetilde{\epsilon}_x = (1 - \theta \beta)x - \theta \epsilon_y$ 

Theorem 1 (Darmois-Skitovich) *Let X*1*,...,X<sup>n</sup> be independent, non-degenerate random variables. If for two linear combinations*

$$
l_1 = a_1 X_1 + \ldots + a_n X_n, \quad a_i \neq 0
$$
  

$$
l_2 = b_1 X_1 + \ldots + b_n X_n, \quad b_i \neq 0
$$

*are independent, then each X<sup>i</sup> is normally distributed.*







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### Bivariate Linear Gaussian case Figure 1 illustrates the basic identifiability principle for the two-variable model. Denoting the two

$$
x = \epsilon_x
$$
  

$$
y = x + \epsilon_y
$$

 $x = \epsilon_x$   $y = \epsilon_y$   $y = \epsilon_y$   $y = \epsilon_y$   $z = \epsilon_x$  and *n*  $z = \epsilon_y$  and *n* are **True model**  $x = \epsilon_x$   $\epsilon_x, \epsilon_y \sim$  indep. Gaussian









### Bivariate Linear Gaussian case Figure 1 illustrates the basic identifiability principle for the two-variable model. Denoting the two



**True model**  $x = \epsilon_x$   $\epsilon_x, \epsilon_y \sim$  indep. Gaussian



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- If  $f_j(.)$  is linear, then non-Gaussian errors are required for identifiability
- $\blacktriangleright$  What if the errors are Gaussian, but  $f_j(.)$  is non-linear?
- More generally, under what circumstances is the causal structure represented by this class of models identifiable?









(true) model

a detailed explanation of (a)–(f). (g) shows an example of a joint density *p*(*x, y*) generated by a causal model *x*  $\alpha$  is non-linear, the support  $\alpha$  is non-linear, the support of the density  $\alpha$  is non-linear, the support of the densities  $\alpha$  is non-linear, the support of the densities  $\alpha$  is non-linear support o (graphics from Hoyer et al. 2009)









causal model *x*  $\alpha$  is non-linear, the support  $\alpha$  is non-linear, the support of the density  $\alpha$  is non-linear, the support of the densities  $\alpha$  is non-linear, the support of the densities  $\alpha$  is non-linear support o (graphics from Hoyer et al. 2009)

**Hoyer Condition (HC)**: *Technical condition on the relation between the function, the noise distribution and the parent distribution that, if satisfied, permits a backward model.*

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	- extension to discrete additive noise models

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- If the function is **linear**, but the error terms **non-Gaussian**, then one can't fit a linear backwards model (Lingam), but there are cases where **one can fit a non-linear backwards model**




• how to integrate data from experiments?

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tier orderings

• how to integrate data from experiments?





$$
\begin{bmatrix} x \\ y \end{bmatrix} \succ \begin{bmatrix} z \\ w \end{bmatrix}
$$

tier orderings



"priors"

• how to integrate data from experiments?



• how to include background knowledge?





tier orderings





• specific search space restrictions

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㱺

 $l_{2}$ 

*x*  $\rightarrow$   $\rightarrow$   $\rightarrow$ 





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# High-Level

## High-Level

#### data sample







High-Level











• **Formulate the independence constraints in propositional logic**



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- **Encode the constraints into one formula.**

$$
x \perp y \iff \neg A \land \neg B \dots
$$

$$
A = 'x \rightarrow y \text{ is present'}
$$

$$
\neg A \land \neg B \land \neg (C \land D) \land \neg \dots
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$$

$$
A = false
$$
  

$$
B = false
$$
  

$$
\Leftrightarrow
$$

*...*

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 $\iff$ *x y z A* = *f alse B* = *f alse*

■ very general setting (allows for cycles and latents) and trivially complete

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■ very general setting (allows for cycles and latents) and trivially complete

*...*

➡ **BUT**: erroneous test results induce conflicting constraints: UNsatisfiable

Conflicts and Errors

• **Statistical independence tests produce errors** 

➡ **Conflict: no graph can produce the set of constraints**



Conflicts and Errors

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➡ **Conflict: no graph can produce the set of constraints**





 $\boldsymbol{x}$ 

*z*

Conflicts and Errors

Sridhar talk

*x y*

*z*

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Constraint Satisfaction Approach

• **INPUT: (in)dependence constraints weighted according to reliability** 

$$
\min_{G} \sum_{k \,:\, \text{constraint } k \text{ is not satisfied by } G}
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#### **What are suitable weights?**

# Weighting Schemes

- Constant weights
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## Weighting Schemes

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	- only treat rejections of the null-hypothesis as hard constraints, in line with classical statistics
	- give dependences infinite weight, maximize the independences (unit weight) in light of these dependences
- Log weights
	- obtain the probability of an (in)dependence and weigh it according to the log of the probability
	- Model selection with Bayes rule:

 $x \cancel{\perp} y$ <sup>|</sup>*C*  $P(x|C)P(y|x, C)$  $x \perp y$ <sup>*C*</sup>  $P(x|C)P(y|C)$
# Simulation 1: no cycles, no latents, linear Gaussian



- TPR vs. FPR of all d-separation constraints of the true graph for a varying p-value cut-off
- observational data set, 6 observed variables, average degree 2; 500 samples, 200 models, linear Gaussian parameterization

• cPC returns a fully determined output only 58/200 times at its optimum

## Simulation 2: no cycles, but latents



• cFCI only returns unambiguous results 61/200 times at its optimum

## Simulation 3: cycles and latents



[Hyttinen et al. 2014]







 $x \mathop{\perp\!\!\!\!\!\!/\!\!\!\!\!\!\perp} w||x \; z$ 



# (b)  $\iff x \not\perp w||x \not z \quad weight = 0.8$

32



# (b)  $\iff x \not\perp w||x \not z \quad weight = 0.8$









# $x \cancel{\perp} w || x z$  *weight* = 0.8

 $\begin{array}{ccc} \text{for } & (x > z) \wedge (x > w) \ & \wedge (y > z) \wedge (y > w) \end{array}$  $\wedge$   $(y>z)$   $\wedge$   $(y>w)$ 

- specific probabilities for each graph
- soft sparsity constraint
- $\bullet$  …





# **Simulation 4: Scalability**



• up to 10 variables and only a few overlapping data sets for now

[Hyttinen et al. 2014]





## Query:



 $\mathbf{N}$ 





• list the structures in the equivalence class

**(max) SAT-solver ISAL-Solver** 



## Query:

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	- edges, confounders
	- ancestral relations
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- what are the highest scoring equivalence classes?

**SALL** 

➟

**(max) SAT-solver** max) SAT-solver



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- what are the highest scoring equivalence classes?

## Response:

- enumeration of solutions
- "backbone" of the SAT-instance

 $\bullet$  …

# **Computing Causal Effects**



# Computing Causal Effects





• search in the equivalence class over the possible applications of the *do*-calculus rules by *querying* the satisfaction of their dseparation conditions



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*do*-calculus

Rule 1 (insertion/deletion of observations)

 $P(y|do(x), z, w) = P(y|do(x), w)$  if  $Y \perp Z | X, W | X$ 

Rule 2 (action/observation exchange)

 $P(y|do(x), do(z), w) = P(y|do(x), z, w)$  if  $Y \perp I_Z | X, Z, W | X$ 

Rule 3 (insertion/deletion of actions)

 $P(y|do(x), do(z), w) = P(y|do(x), w)$  if  $Y \perp I_Z|X, W||X$ 

[Hyttinen et al. 2015]



• search in the equivalence class over the possible applications of the *do*-calculus rules by *querying* the satisfaction of their dseparation conditions

 $P(y|do(x), z, w) = P(y|do(x), w)$  if  $Y \perp Z | X, W | X$  $P(y|do(x), do(z), w) = P(y|do(x), z, w)$  if  $Y \perp I_Z|X, Z, W||X$  $P(y|do(x), do(z), w) = P(y|do(x), w)$  if  $Y \perp I_Z|X, W||X$ *do*-calculus Rule 1 (insertion/deletion of observations) Rule 2 (action/observation exchange) Rule 3 (insertion/deletion of actions)



High-Level

tting ime series<br>hternal latent structures<br>:tc.

e as ints on graph j<br>J

**(max) SAT-solver**max) SAT-solver

High-Level

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**(max) SAT-solver** max) SAT-solver **QUERY?**



```
Just getting started...
```
- Just getting started…
- application



[Stekhoven et al. 2012]

• application



[Stekhoven et al. 2012]

• multi-scale causal analysis: micro- to macro-variables



[Chalupka et al. 2016]

• application



[Stekhoven et al. 2012]

• multi-scale causal analysis: micro- to macro-variables



#### [Chalupka et al. 2016]

• time-series and dynamics



• application



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• violations of the Markov property: non-causal relations



[Maier et al. 2013]

• application Sokolova talk



#### [Stekhoven et al. 2012]

• multi-scale causal analysis: micro- to macro-variables



#### [Chalupka et al. 2016]



• time-series and dynamics



• violations of the Markov property: non-causal relations



[Maier et al. 2013]

# References

#### **Limitations**

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### <sup>40</sup> Thank you!